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PREDICTION OF STATISTICAL
SYSTEM PERFORMANCE FROM
PARAMETER DISTRIBUTIONS

by

Alvin LaVerne Franson

United States Naval Postgraduate School



THESIS

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June 1969

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Prediction of Statistical System Performance
from
Parameter Distributions

by

Alvin LaVerne Franson
Lieutenant, United States Navy
B. S. E. E. , University of Kansas, 1963

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ABSTRACT

Techniques for extrapolating statistical data on the parameters of individual components to the statistical performance index for an overall system are considered. Two cases are evaluated. (1) The deterministic case in which the system's performance index is known functionally in terms of the system parameters. (2) The non-deterministic case in which only limited data on the performance index and its sensitivity with respect to system parameters is known. Computer programs are developed in both cases for combining given probability density distributions for the parameters into an overall probability density distribution for the system performance index. Theory and programs are developed and verified with a specific numerical example.

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I. INTRODUCTION

The performance of a system, from a statistical point of view, depends upon the nominal values of the parameters of the component parts included in the system and on the statistical distributions of those parameters. Alternately, the response of the component parts to a given signal may be known statistically and it is desired to predict an overall statistical distribution for the system response. Statistical information on the system's response may be interpreted in terms of probability of success or failure (operation or non-operation) for a single system or in terms of the percentage of operational systems that can be expected in an ensemble of systems. Currently three basic methods are being used to obtain performance data for systems; the worst-case method [Mark and Stember 1965], the moment method [Mark and Stember 1965], and the Monte Carlo method [Shreider 1966].

The worst-case method is a non-statistical approach intended to determine whether it is possible, with given parameter tolerance limits, for the systems performance characteristics to fall outside specifications. This method determines nothing of the probability distribution of the performance. Worst-case analysis, although it may be involved, does provide useful engineering insight and is used a great deal in practice. However, its outlook is basically limited.

Statistics are combined with system analysis techniques in the moment method to estimate the probability that performance will remain within specified limits. The actual parameter probability distributions are usually replaced by Gaussian distributions having the same variances. Hence, the moment method has inherent limitations. Parameter probability distribution characteristics such as skewness and peakedness are usually not taken into account.

In the Monte Carlo method a large number of alternate replicas of a system are simulated by mathematical models. Component values are randomly selected and the performance of each replica is determined for its particular set of components. The accuracy of the Monte Carlo method is dependent upon taking a sufficiently large number of replicas and can be both lengthy and costly. Monte Carlo methods are also used extensively in practice although they provide limited engineering information.

The purpose of this thesis is to present a numerical technique for combining statistical data on several independent components into a single distribution function for the system performance. This is accomplished by considering two general cases. The deterministic case, where the functional relationships between the system parameters and the output are known functionally; and the non-deterministic case, where these functional relationships are unknown but measured discrete data is available. In both cases the actual parameter probability distributions are used in conjunction with numerical integration

techniques to obtain the system distribution function. One major advantage of the approach is that the known parameter distributions are quite arbitrary and may be histograms obtained from the results of experiments on individual components.

II. DETERMINISTIC CASE

The object of the deterministic case is to investigate the distribution of a system output when the output has a known mathematical relationship to several internal parameters. The basic assumptions are that all of the internal parameters are independent and that the probability distributions of the internal parameters, although arbitrary, are known mathematically or given in terms of histograms.

To commence the investigation, the general equations for various combinations of parameter distributions which may be expected in a typical system performance index must be obtained. These will basically consist of operations on the distributions of random variables which are interrelated by addition, subtraction, multiplication, and division. The basic mathematics for these operations are well known in the literature, at least in general form [Beckmann 1967].

Consider two independent random variables: X , assuming values x with a known probability density $p_X(x)$, and Y , assuming values y with a known probability density $p_Y(y)$. Also consider the functions $U = U(X, Y)$ and $V = V(X, Y)$ and their inverse functions $X = X(U, V)$ and $Y = Y(U, V)$. Then, when X is near x and Y is near y , U and V must be near $u = u(x, y)$ and $v = v(x, y)$. Hence,

$$p_{XY}(x, y) \, dx dy = p_{UV}(u, v) \, du dv \quad (1)$$

Substituting the inverse functions and the required probability densities, we obtain

$$p_{UV}(u, v) = p_X [x(u, v)] p_Y [y(u, v)] \left| \frac{\partial(x, y)}{\partial(u, v)} \right| \quad (2)$$

where $p_{UV}(u, v)$ is the joint probability density of U and V , $p_X[x(u, v)]$ is the probability density of X , $p_Y[y(u, v)]$ is the probability density of Y , $x(u, v)$ and $y(u, v)$ are the inverse functions, and $\left| \frac{\partial(x, y)}{\partial(u, v)} \right|$ is the Jacobian, given by

$$\frac{\partial(x, y)}{\partial(u, v)} = \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} \quad (3)$$

When X and Y are related by a single functional the foregoing can be modified by letting $U = U(X, Y)$ and $V = X$. (4)

Then,

$$\frac{\partial(x, y)}{\partial(u, v)} = \begin{vmatrix} \frac{\partial x}{\partial u} & 1 \\ \frac{\partial y}{\partial u} & 0 \end{vmatrix} = - \frac{\partial y}{\partial u} \quad (5)$$

Equation (2) then becomes

$$p_{UX}(u, x) = p_X(x) p_Y [y(u, x)] \left| \frac{\partial y}{\partial u} \right| \quad (6)$$

Since, in general, $p_U(u) = \int_{-\infty}^{\infty} p_{UX}(u, x) dx$, equation (6) may be written as

$$p_U(u) = \int_{-\infty}^{\infty} p_X(x) p_Y [y(u, x)] \left| \frac{\partial y}{\partial u} \right| dx \quad (7)$$

Using equation (7), the specific formulas of Table I have been compiled for the functional relationships and operations listed, which were chosen on the basis of what might be expected in engineering problems.

The technique outlined here consists of the successive application of the formulas of Table I as required, with the required integration performed numerically. Various types of numerical integration are possible; rectangular integration, trapezoidal integration, or Simpson's Rule. These are summarized as follows:

Rectangular integration [Stanton 1961],

$$\int_{x_1}^{x_n} f(x) dx = \Delta x \sum_{i=1}^n f(x_i) \quad (8)$$

Trapezoidal integration [Stiefel 1963],

$$\int_{x_1}^{x_n} f(x) dx = \Delta x \left\{ \frac{1}{2} [f(x_1) + f(x_n)] + \sum_{i=2}^{n-1} f(x_i) \right\} \quad (9)$$

Simpson's Rule [Stanton 1961],

$$\int_{x_0}^{x_0+n\Delta x} f(x) dx = \frac{\Delta x}{3} \sum_{i=0}^n C_i f(x_i) \quad (10)$$

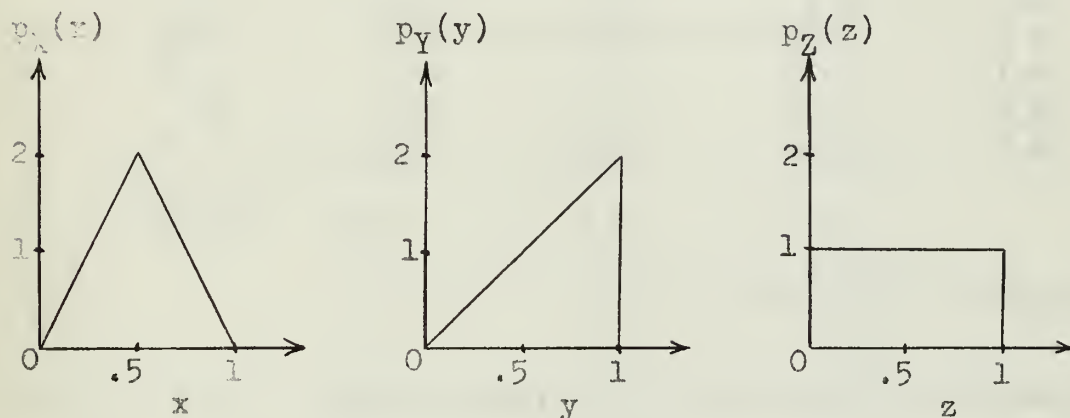
where $C = 1, 4, 2, \dots, 2, 4, 1$

Although the specific technique of numerical integration is of prime importance in so far as accuracy and computer programing and time are concerned, the exact procedure used is a detail which is best dictated by the particular problem to which these techniques are to be applied.

This method may be best illustrated by considering the following example. Consider that:

$$H = \sum_{i=1}^n x_i^2 + b + Z \quad (11)$$

where a and b are constants and X , Y , and Z are independent random variables with known (measured) probability densities. Consider that $p_X(x)$, $p_Y(y)$, and $p_Z(z)$ are given as follows:



(a) Consider first the combination $U = XY$, with $\Delta x = 0.1$.

Using rectangular integration for simplicity, equation (3) of Table I

becomes, for $u = u_1$:

$$p_U(u_1) = \int_{-\infty}^{\infty} p_X(x) p_Y(u/x) \frac{dx}{|x|} \approx \Delta x \sum_{i=1}^9 \frac{p_X(x_i)}{x_i} p_Y\left(\frac{u_i}{x_i}\right) \quad (12)$$

When $u = u_1 = 0.09$

x_i	u_i/x_i	$\frac{p_X(x_i)}{x_i}$	$p_Y(u_i/x_i)$	$\frac{p_X(x_i)}{x_i} p_Y(u_i/x_i)$
0.1	0.9	4.0	1.8	7.2
0.2	0.45	4.0	0.9	3.6
0.3	0.3	4.0	0.6	2.4
0.4	0.22	4.0	0.43	1.72
0.5	0.18	4.0	0.37	1.48
0.6	0.15	2.6	0.30	0.78
0.7	0.13	1.7	0.26	0.44
0.8	0.11	1.0	0.21	0.21
0.9	0.1	0.44	0.20	0.088

sum = 17.918

$$p_U(u_1) = (0.1)(17.918) = 1.79$$

when $u=u_2=0.2$

x_i	u_2/x_i	$\frac{p_x(x_i)}{x_i}$	$p_Y(u_2/x_i)$	$\frac{p_x(x_i)}{x_i} p_Y(u_2/x_i)$
0.1	2.0	4.0	0.0	0.0
0.2	1.0	4.0	2.0	8.0
0.3	0.7	4.0	1.4	5.6
0.4	0.5	4.0	1.0	4.0
0.5	0.4	4.0	0.8	3.2
0.6	0.33	2.6	0.67	1.74
0.7	0.3	1.7	0.6	1.02
0.8	0.25	1.0	0.5	0.5
0.9	0.22	0.44	0.42	0.185
sum =				24.245

$$p_U(u_2) = (0.1)(24.245) = 2.4$$

The pattern for evaluating $p_U(u)$ is repeated for $u = .3, .4, \dots, 1.0$ yielding the curve of Figure 1.

(b) Next consider the function $V = \frac{1}{U} = \frac{1}{XY}$. Here equation (7) of Table I may be applied.

$$p_V(v) = -\frac{1}{v^2} p_U(1/v) = v^2 p_U(u) \quad (13)$$

u	v	$p_U(u)$	u^2	$p_V(v)$
0.1	10.0	1.8	0.01	0.018
0.2	5.0	2.42	0.04	0.097
0.3	3.33	2.39	0.09	0.215
0.4	2.5	2.11	0.16	0.338
0.5	2.0	1.54	0.25	0.385
0.6	1.66	1.02	0.36	0.364
0.7	1.43	0.55	0.49	0.270
0.8	1.25	0.28	0.64	0.179
0.9	1.11	0.05	0.81	0.040

$p_V(v)$ is plotted in Figure 2.

(c) The function W may now be calculated using equation (8)

$$\text{of Table I, where } W = aV + b = \frac{a}{XY} + b \quad (14)$$

and

$$p_W(w) = \frac{1}{|a|} p_V\left(\frac{w-b}{a}\right) = \frac{1}{|a|} p_V(v) \quad (15)$$

If $a = 2$ and $b = 1$, the following result is obtained.

<u>V</u>	<u>W = aV + b</u>	<u>$p_V(v)$</u>	<u>$p_V(v)/2$</u>
1	3	0.0	0.0
2	5	0.385	0.1925
3	7	0.25	0.125
4	9	0.15	0.075
5	11	0.10	0.050
6	13	0.06	0.030
7	15	0.04	0.020
8	17	0.03	0.015
9	19	0.02	0.010
10	21	0.01	0.005

The result is plotted in Figure 3.

(d) Since, from equations (11) and (14),

$$H = W + Z \quad (16)$$

it follows from equation (1) of Table I that $p_H(h)$ is given by the convolution of $p_W(w)$ and $p_Z(z)$. This operation may also be carried out numerically and for the given $p_Z(z)$ may be expressed as

$$\left. \begin{aligned} p_H(h) &= \int_0^h p_W(w) dw & \text{for } 0 \leq h \leq 1 \\ p_H(h) &= \int_{h-1}^h p_W(w) dw & \text{for } h > 1 \end{aligned} \right\} \quad (17)$$

These integrals may also be carried out numerically yielding the results plotted in Figure 4.

To work more complicated problems than the previous example requires a great amount of mathematical calculations. Thus, when using numerical integration techniques the method lends itself nicely to the digital computer. Figures 5 through 9 are the computer flow graphs corresponding to some of the equations of Table I. The computer program for any mathematical function can then be derived from the appropriate combinations of these basic flow graphs.

To illustrate a more complex problem consider the following example:

given

$$\theta = 1285 \left(\frac{1}{A} - \frac{1}{B} \right) + W \quad (18)$$

where

$$A = B_{3a} E_{6a}, \quad B = B_{3b} B_{6b}$$

and

$$B_{3a} = \frac{B_{3b}}{1 + B_{3b}(2.16 \times 10^{-2})}$$

$$B_{6a} = \frac{B_{6b}}{1 + B_{6b}(0.63 \times 10^{-2})}$$

B_{3b} is a normal distribution; mean = 130, standard deviation = 10

B_{6b} is a normal distribution; mean = 65, standard deviation = 12

W is a uniform distribution from $w = -4/7$ to $w = +4/7$

Theta is system output parameter and the variables B_{3b} , B_{6b} , and

W are independent.

Since the variables are independent, equation (18) can first be simplified by ordinary mathematical manipulation.

$$\text{Let } X = B_{3b}, Y = B_{6b}, k_1 = 2.16 \times 10^{-2}, k_2 = 0.63 \times 10^{-2}$$

$$\text{then } B = XY$$

$$A = \frac{XY}{(1 + k_1 X)(1 + k_2 Y)}$$

$$\frac{1}{A} - \frac{1}{B} = k_1 k_2 + \frac{k_1}{Y} + \frac{k_2}{X}$$

$$\text{and } \theta = 0.175 + \frac{27.8}{Y} + \frac{8.1}{X} + W \quad (19)$$

Now, utilizing the basic methods and flow graphs, the problem can be programed on the digital computer and the probability density of θ readily obtained. The program and solution are contained in Appendix I.

III. NON-DETERMINISTIC CASE

In many practical problems the functional relationship between variables is unknown and it cannot be expressed simply by formula. It is generally available only as numerical data from a computer run on a mathematical model of the system under study. Two cases are considered here. First, the case where the system response of interest, $U(X)$, and its sensitivity coefficients, $U'(X)$, are available at discrete values of a single variable, X .

$$U(X) = [u(x_1), u(x_2), - - -, u(x_n)] \quad (20)$$

$$U'(X) = [u'(x_1), u'(x_2), - - -, u'(x_n)] \quad (21)$$

Second, the case where the system response of interest, $U(X, Y)$, and its sensitivity coefficients, $U'(X, Y)$, are available at discrete values of two independent variables, X and Y .

$$U(X, Y) = \begin{vmatrix} u_{11} & u_{12} & - & - & - & u_{1n} \\ u_{21} & u_{22} & - & - & - & u_{2n} \\ - & - & & & & - \\ - & - & & & & - \\ u_{m1} & u_{m2} & - & - & - & u_{mn} \end{vmatrix} \quad (22)$$

where

$$u_{ij} = u(x_i, y_j)$$

and

$$U'(X, Y) = \begin{vmatrix} u'_{11} & u'_{12} & - & - & u'_{1n} \\ u'_{21} & u'_{22} & - & - & u'_{2n} \\ - & - & & & - \\ - & - & & & - \\ u'_{m1} & u'_{m2} & - & - & u'_{mn} \end{vmatrix} \quad (23)$$

where

$$u'_{ij} = \frac{\partial u(x_i, y_j)}{\partial x_i}$$

In both cases the problem reduces to determining $p_U(u)$ from the known distributions $p_X(x)$, $p_Y(y)$ and the data of equations (20) and (21), or (22) and (23).

In the first case, consider that the probability density, $p_X(x)$, and the cumulative distribution,

$$P_X(x_j) = \int_{-\infty}^{x_j} p_X(x) dx \quad (24)$$

are known. Then at each data point of $u(x)$,

$$P_U[u(x_j)] = P_X(x_j) \quad (25)$$

That is, the probability that $u \leq u(x_j)$ is equal the probability that $x \leq x_j$. Equation (25) gives discrete points on the cumulative distribution, $P_U(u)$, for each value of the parameter x for which $u(x)$ has been measured.

It is also well known [Beckman 1967] that

$$p_U(u) = \frac{\Delta P_U(u)}{\Delta u} = \frac{p_X(x)}{\left| \frac{du}{dx} \right|} \quad (26)$$

Thus the slope of the cumulative distribution curve, $\frac{P_U(u_i)}{\Delta y}$, at each data point is equal to the known probability density, $p_X(x_j)$, divided by the measured sensitivity function, $\left| \frac{\partial x}{\partial y} \right|$, at the data point. The conclusion is that the cumulative distribution, $P_U(u)$, and its slope may be calculated readily from the measured response and sensitivity data. This curve is therefore evaluated as a piece-wise linear approximation.

When U is a function of two variables, X and Y , the foregoing cumulative distribution curve can be evaluated for each value of the second variable, Y . Thus a family of such curves or data points denoted by $P_U(u, y_j)$ is available. For a specific value of $u = u_i$ each of these curves gives the probability that $u \leq u_i$ when $y = y_j$. The total probability that $u \leq u_i$ depends upon $p_Y(y)\Delta y$ and is given by the equation:

$$P_U(u_i) = \Delta y \sum_{j=1}^m P_U(u_i, y_j) p_Y(y_j) \quad (27)$$

where $P_U(u, y_j)$ is the cumulative distribution evaluated with y as a constant and x as a single variable as discussed previously.

This result can be proven by considering the following (see equation (7)):

$$P_U(u) = \int_{-\infty}^{\infty} p_Y(y) p_X[x(u, y)] \left| \frac{\partial x}{\partial y} \right| dy \quad (28)$$

where $u = u(x, y)$. $\frac{\partial x}{\partial y}$ is the sensitivity function with respect to the parameter x and in equation (28) must be expressed in terms of u and y . It is convenient to denote

$$\frac{\partial u}{\partial x} = u_X(u, y) \quad (29)$$

and then equation (28) becomes

$$p_U(u) = \int_{-\infty}^{\infty} p_Y(y) p_X[x(u, y)] \frac{dy}{u_X(u, y)} \quad (30)$$

We also know

$$P_U(u) = \int_{-\infty}^u p_U(u) du \quad (31)$$

Substituting equation (30) into equation (31) and inter-changing the order of integration yields

$$P_U(u) = \int_{-\infty}^{\infty} p_Y(y) dy \int_{-\infty}^u p_X[x(u, y)] \frac{du}{u_X(u, y)} \quad (32)$$

Now

$$\int_{-\infty}^u p_X[x(u, y)] \frac{du}{u_X(u, y)} = \int_{-\infty}^u p_U(u, y) du = P_{UY}(u, y) \quad (33)$$

Since, if Y is considered a constant and X as the random variable,

$u = u(x, y)$,

$$p_U(u, y) = \frac{p_X[x(x, y)]}{\frac{\partial u}{\partial x}} = \frac{p_X[x(x, y)]}{u_X(u, y)} \quad (34)$$

Substituting equation (33) into equation (32) yields:

$$P_U(u) = \int_{-\infty}^{\infty} p_Y(y) P_U(u, y) dy \quad (35)$$

In discrete form, equation (35) becomes equation (27) when rectangular integration is used. However, more accurate integration techniques may also be applied to equation (35).

To demonstrate the application of this non-deterministic method consider part of the previous example of equation (18). It is now assumed that the functional relationship between the output and the parameters is unknown. However, we do know the probability densities of the parameters x and y and have certain measured data which consists of measured values of x , y , and \emptyset and the corresponding sensitivities for each measured \emptyset . The data is presented in Table II. Following the procedure outlined, the problem can be programed on the digital computer and the probability density of \emptyset readily obtained. The program and solution are contained in Appendix II. The results can be compared with those obtained using the deterministic formulas. Noting, of course, that the solution to the deterministic problem, contained in the Appendix, includes the addition of the uniform distribution $W(w)$ and the solution to the non-deterministic problem does not.

IV. CONCLUSIONS

The numerical technique presented here has distinct advantages over the worst-case, the moment, and the Monte Carlo methods. By utilizing the actual parameter probability distributions the results obtained are more accurate than those of the moment method and present a great deal more information than the worst-case method. The accuracy obtained is equivalent to that of a large number of Monte Carlo replicas. Thus the saving in time over the Monte Carlo method is obvious. In general, the numerical method presented here can obtain a more accurate solution in less time than any of the other current methods. However, certain limitations are inherent in the basic assumptions used to formulate the method. Since independent parameters are required, a great number of practical problems can not be handled by this method as it is presented here. By extending the mathematical derivations to include the correlation of parameters, the method could be used for systems in which the parameters are not independent.

Table I
Functional Formulas

$$\begin{aligned} (1) \quad U = X + Y \quad p_U(u) &= \int_{-\infty}^{\infty} p_X(x) p_Y(u-x) dx \\ &= \int_{-\infty}^{\infty} p_X(u-y) p_Y(y) dy \end{aligned}$$

$$\begin{aligned} (2) \quad U = X - Y \quad p_U(u) &= \int_{-\infty}^{\infty} p_X(x) p_Y(x-u) dx \\ &= \int_{-\infty}^{\infty} p_X(u+y) p_Y(y) dy \end{aligned}$$

$$(3) \quad U = XY \quad p_U(u) = \int_{-\infty}^{\infty} p_X(x) p_Y(u/x) \frac{dx}{|x|}$$

$$\begin{aligned} (4) \quad U = X/Y \quad p_U(u) &= \frac{1}{u^2} \int_{-\infty}^{\infty} p_X(x) p_Y(x/u) |x| dx \\ &= \int_{-\infty}^{\infty} p_X(uy) p_Y(y) |y| dy \end{aligned}$$

$$(5) \quad U = X^2 \quad p_U(u) = \frac{p_X(\sqrt{u}) + p_X(-\sqrt{u})}{2\sqrt{u}} \quad u \geq 0$$

$$(6) \quad U = |X| \quad p_U(u) = p_X(x) + p_X(-x) \quad x \geq 0$$

$$(7) \quad U = a/X \quad p_U(u) = \frac{a}{u^2} p_X(a/u)$$

$$(8) \quad U = aX + b \quad p_U(u) = \frac{1}{a} p_X\left(\frac{u-b}{a}\right)$$

(a and b are constants)

Table II

Sample Data

$\chi \backslash y$	41	53	65	77	89
110	0.9266	0.7731	0.6761	0.6096	0.5616
120	0.9204	0.7669	0.6699	0.6034	0.5554
130	0.9152	0.7617	0.6647	0.5982	0.5502
140	0.9108	0.7573	0.6603	0.5938	0.5458
150	0.9069	0.7534	0.6564	0.5899	0.5419

 θ

y	41	53	65	77	89
$\left \frac{\partial \theta}{\partial y} \right $	0.0165	0.00988	0.00658	0.00469	0.00352

χ	110	120	130	140	150
$\left \frac{\partial \theta}{\partial \chi} \right $	0.000670	0.000561	0.000478	0.000413	0.000359

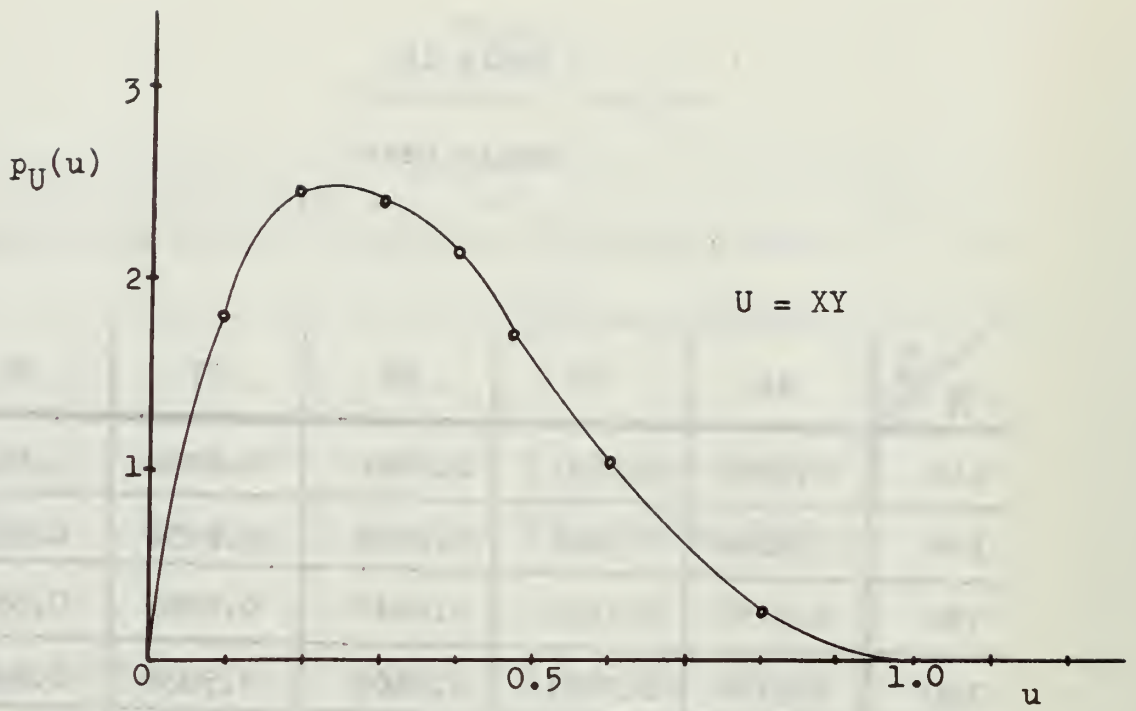


Figure 1

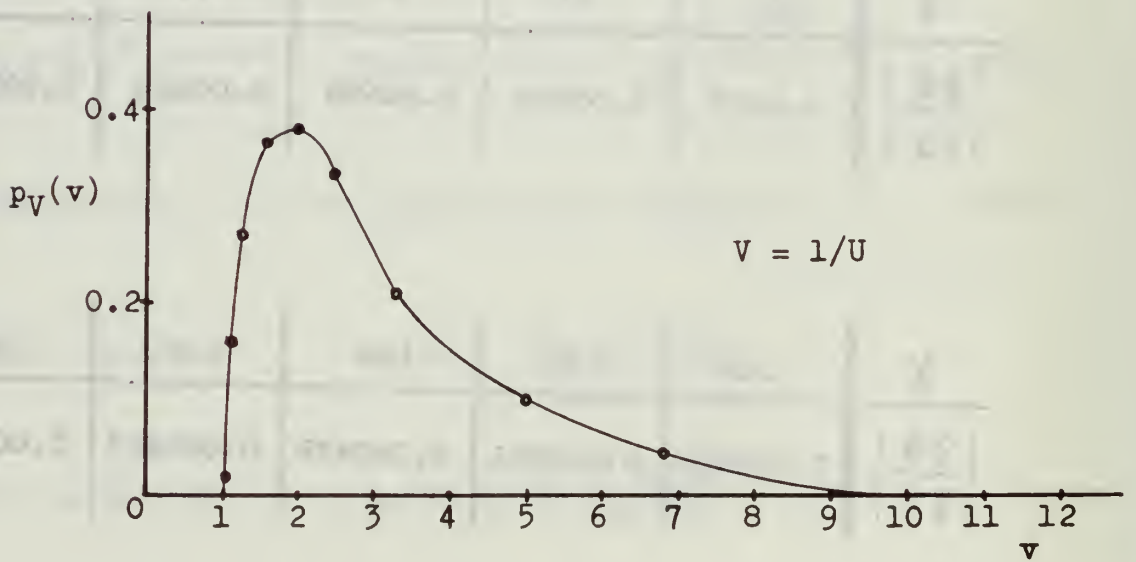


Figure 2

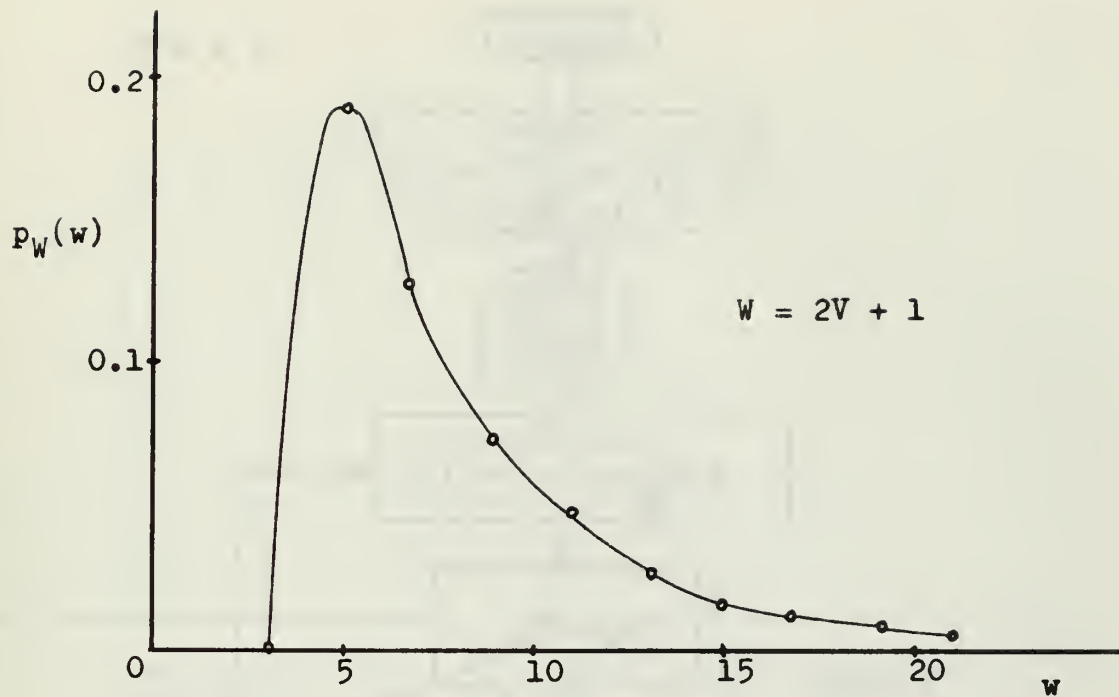


Figure 3

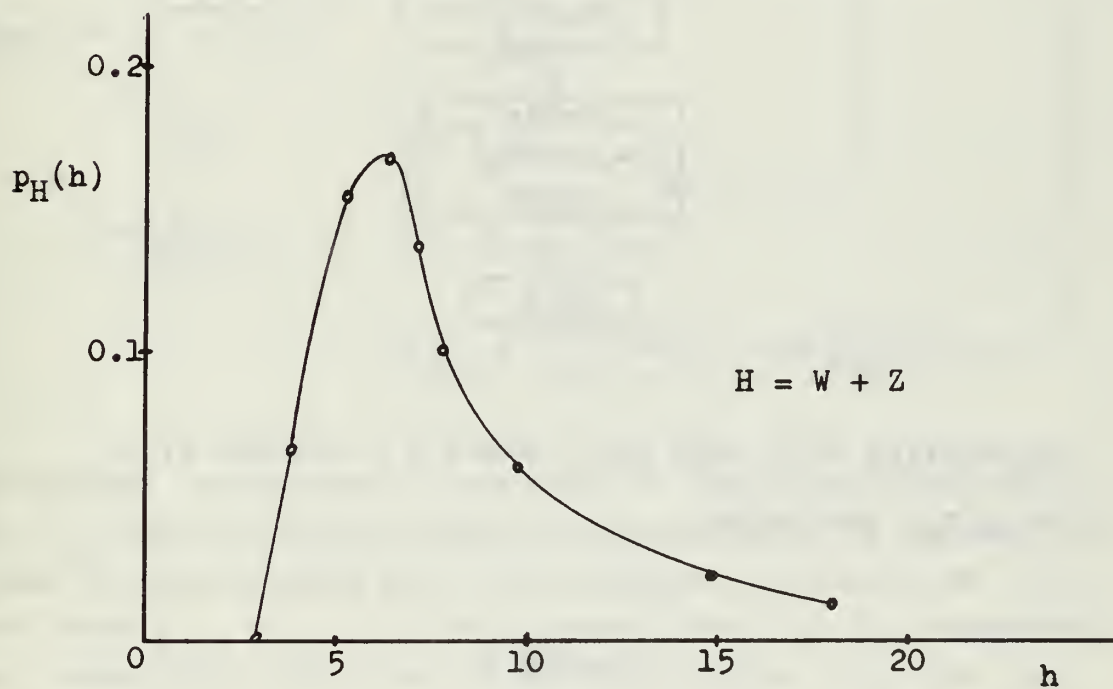
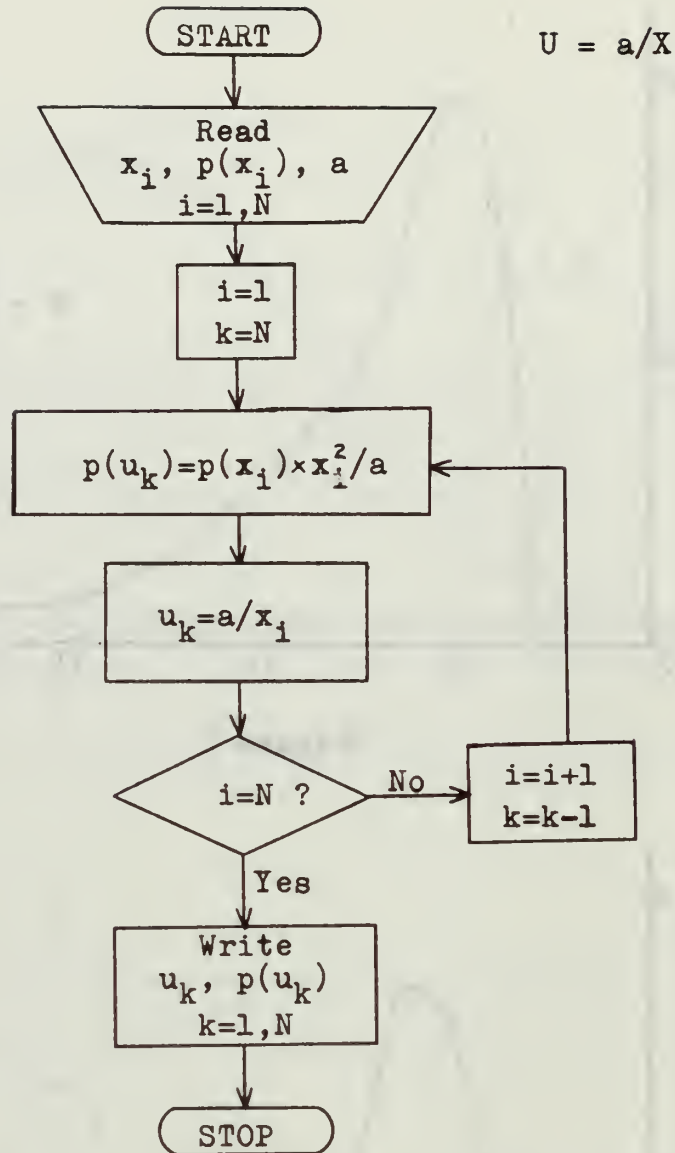
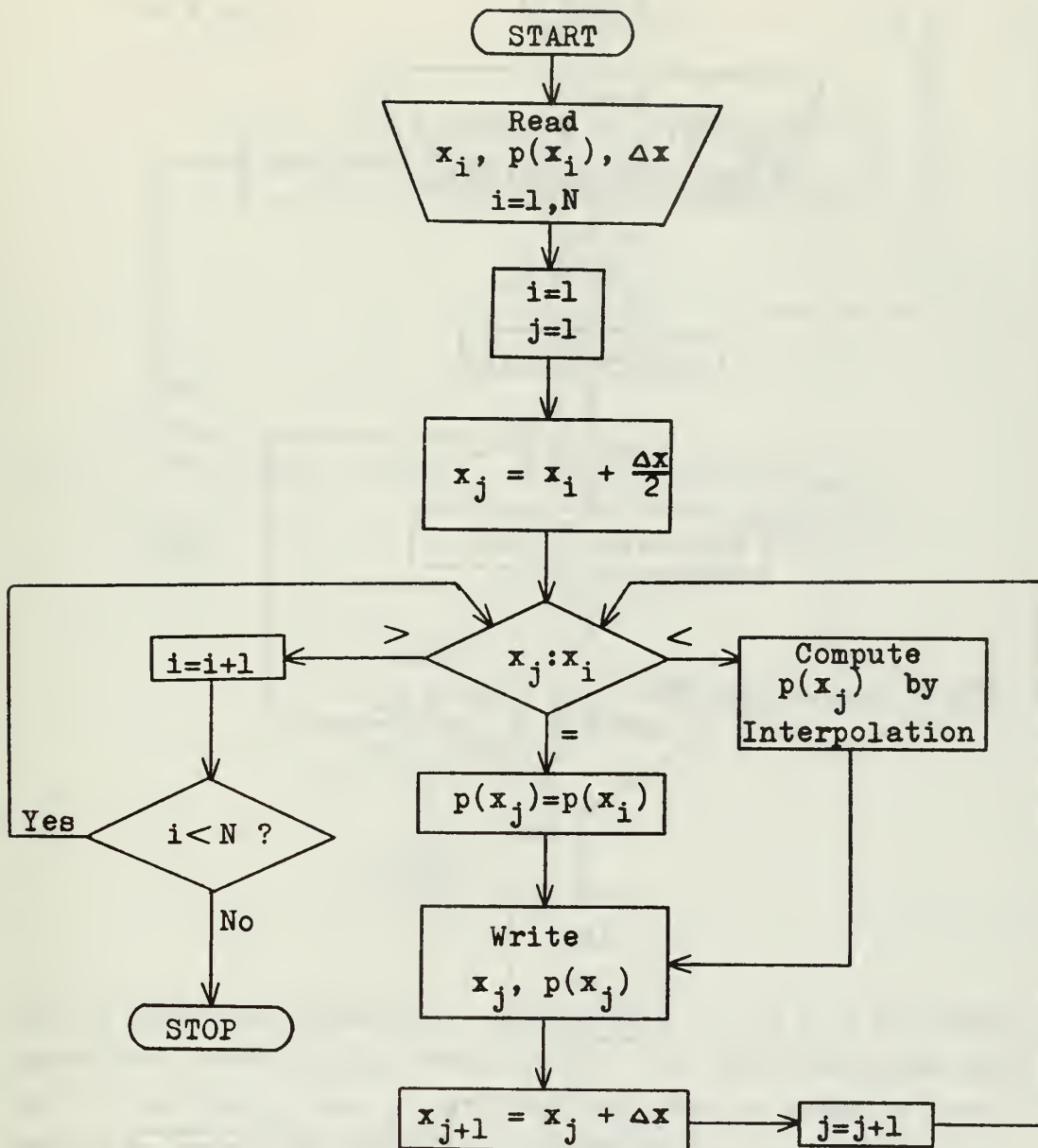


Figure 4



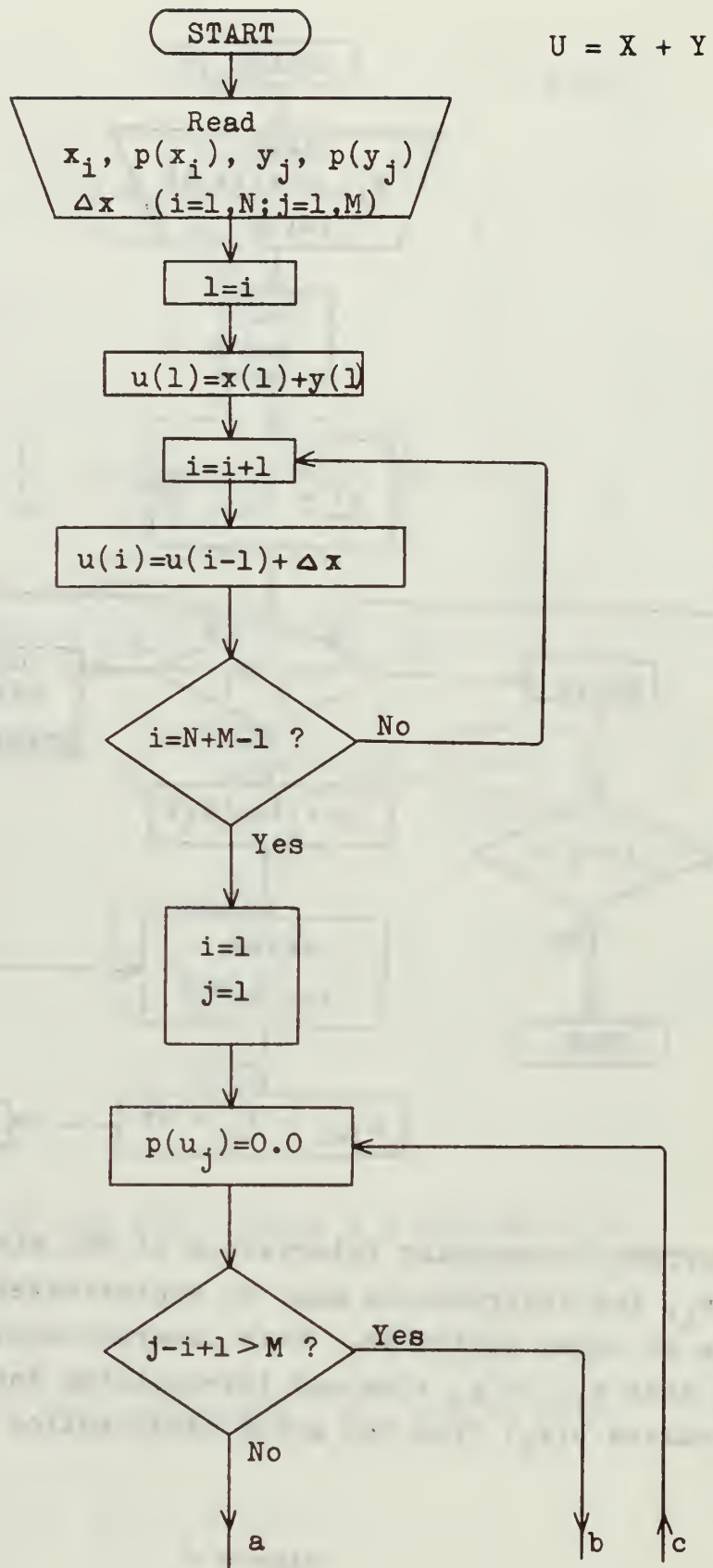
Calculation of u_k and $p(u_k)$ where $U = a/X$ and a , x_i , and $p(x_i)$ are known.

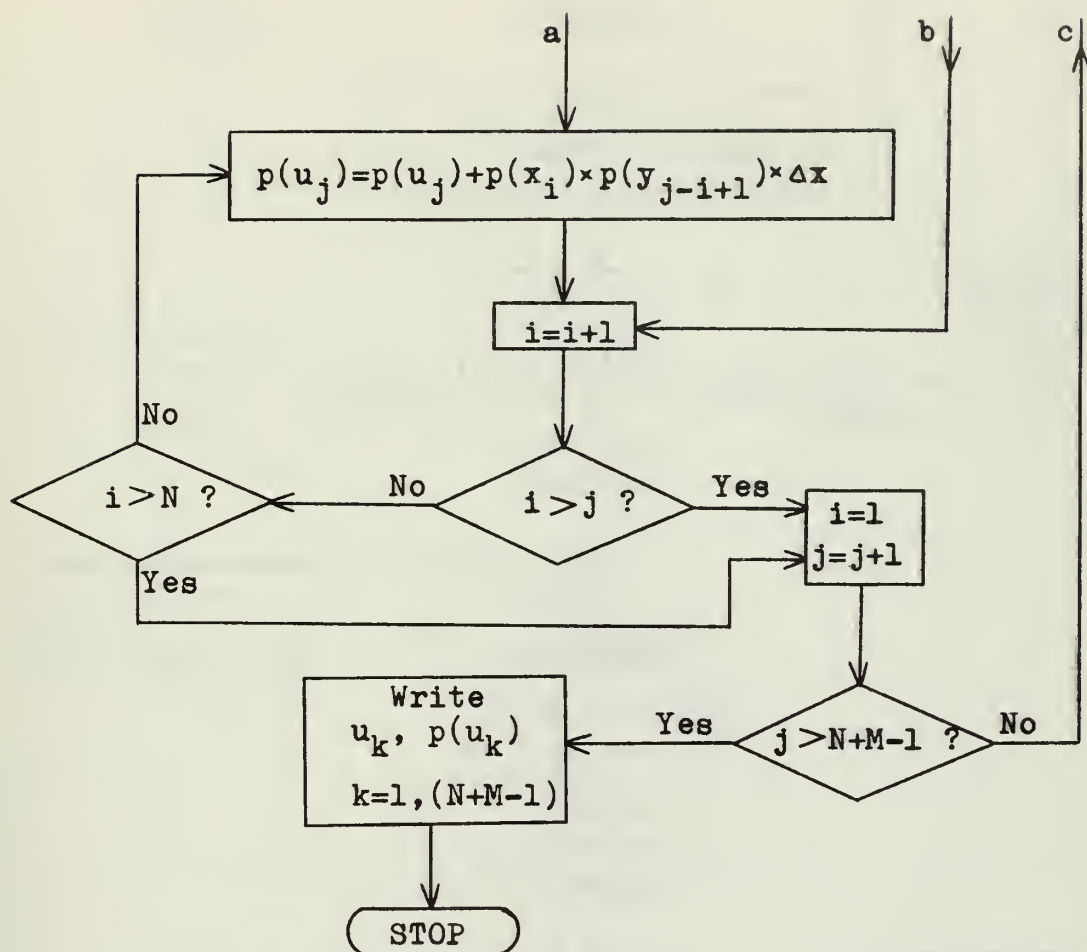
Figure 5



To perform rectangular integration of the distribution $p(x_i)$ vs. x_i , the distribution must be approximated by rectangular areas of equal widths Δx . This program calculates the x_j 's such that $x_{j+1} = x_j + \Delta x$ and interpolates for the corresponding values $p(x_j)$ from the given distribution $p(x_i)$ vs. x_i .

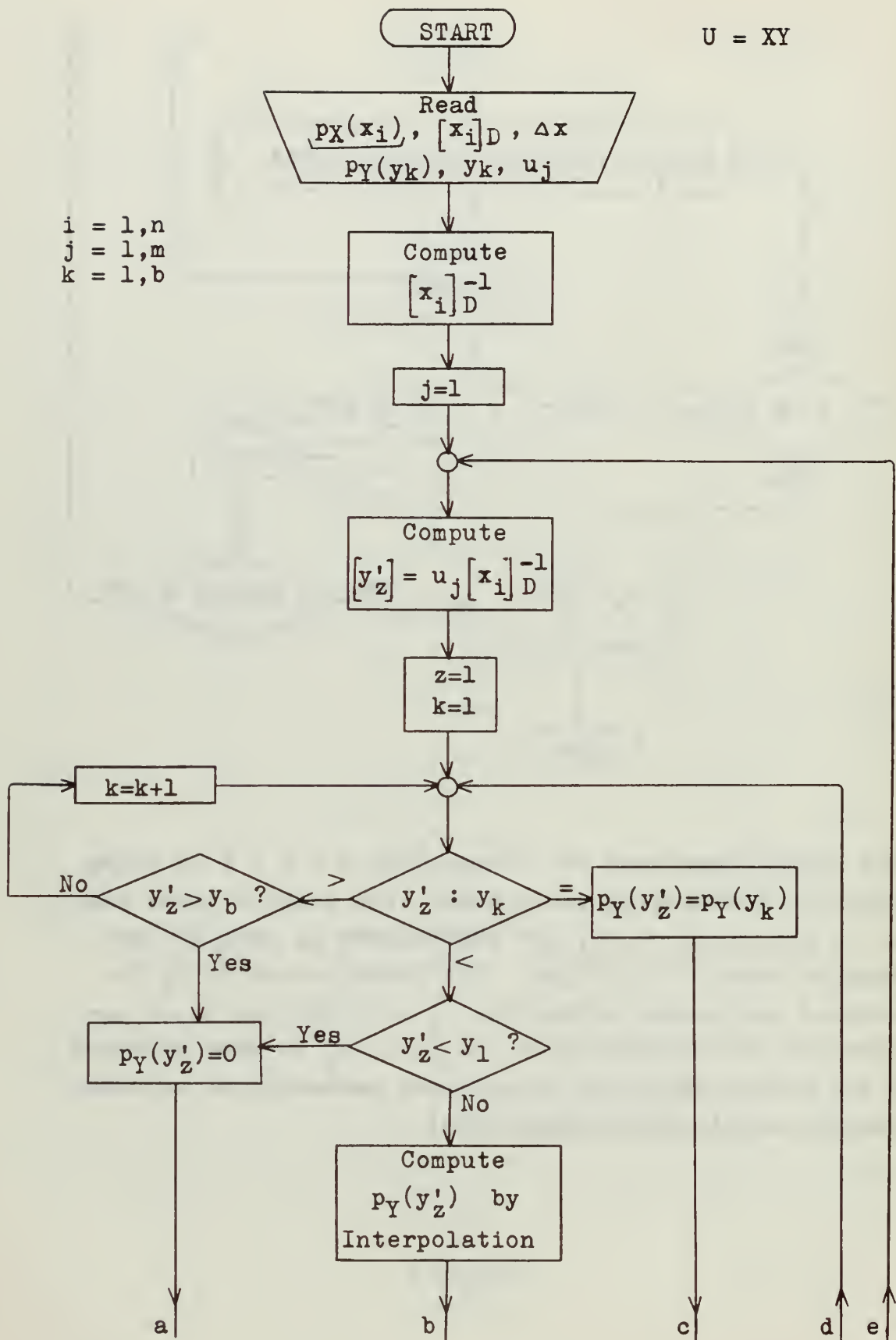
Figure 6





This program performs the convolution $U = X + Y$ by using graphical convolution techniques. The distributions $p(x_i)$ vs. x_i and $p(y_j)$ vs. y_j are represented by sets of rectangular areas of width Δx . The distribution of x_i is reversed and passed across the y_j distribution in Δx increments. The distribution $p(u_k)$ vs. u_k is then obtained by the normal graphical convolution mathematical calculations [Bracewell 1965, Healy 1969].

Figure 7



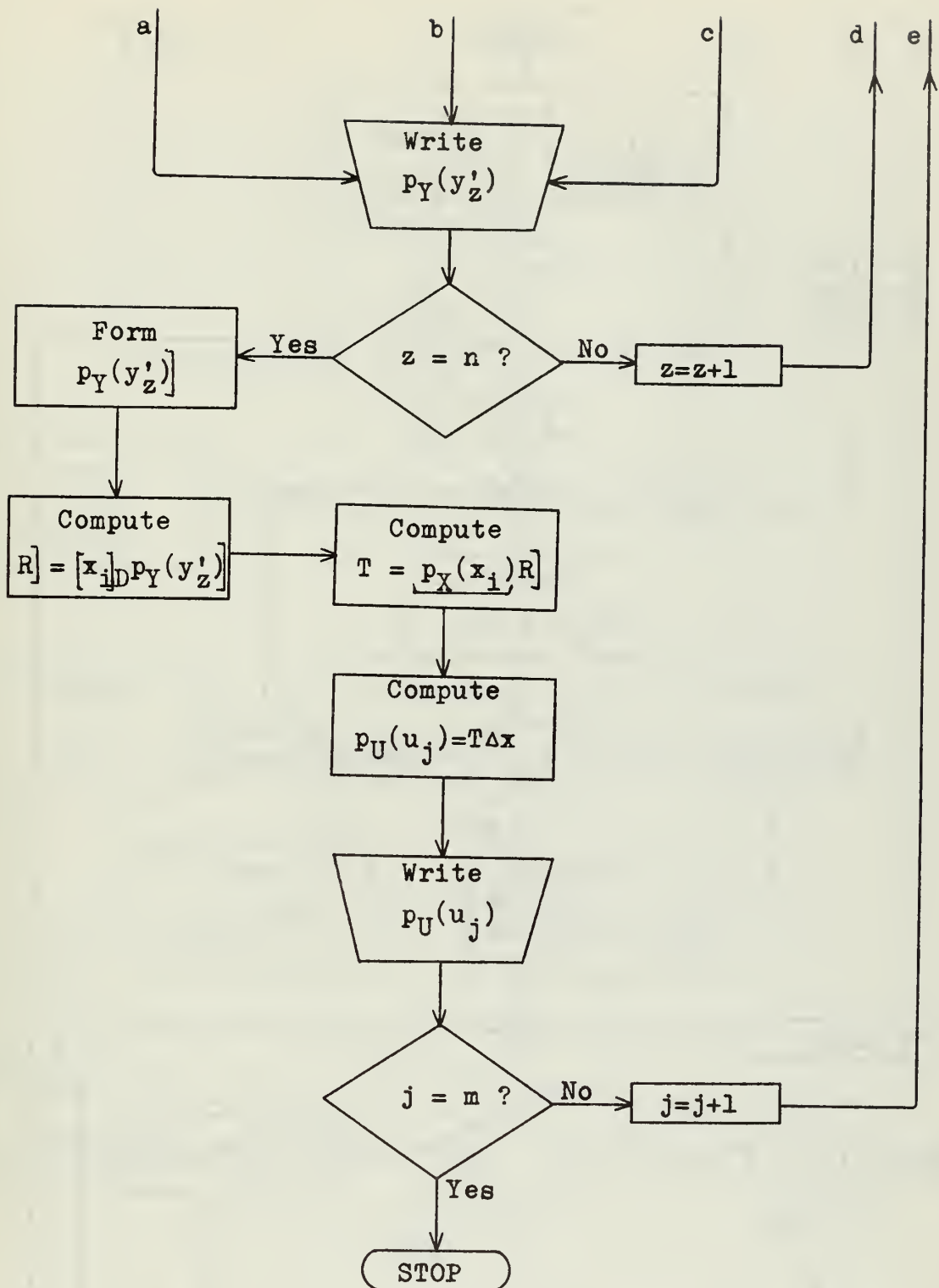
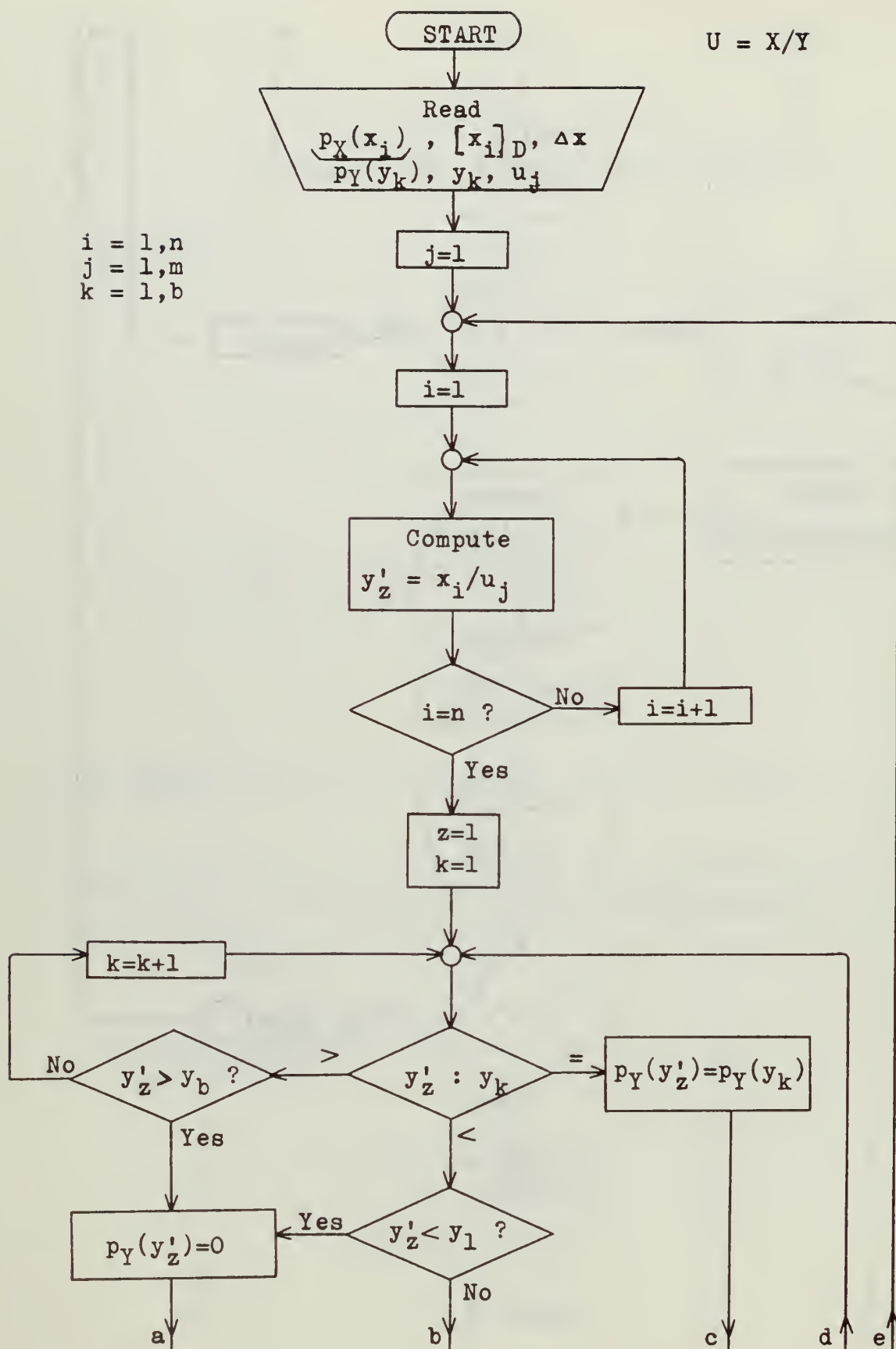


Figure 8



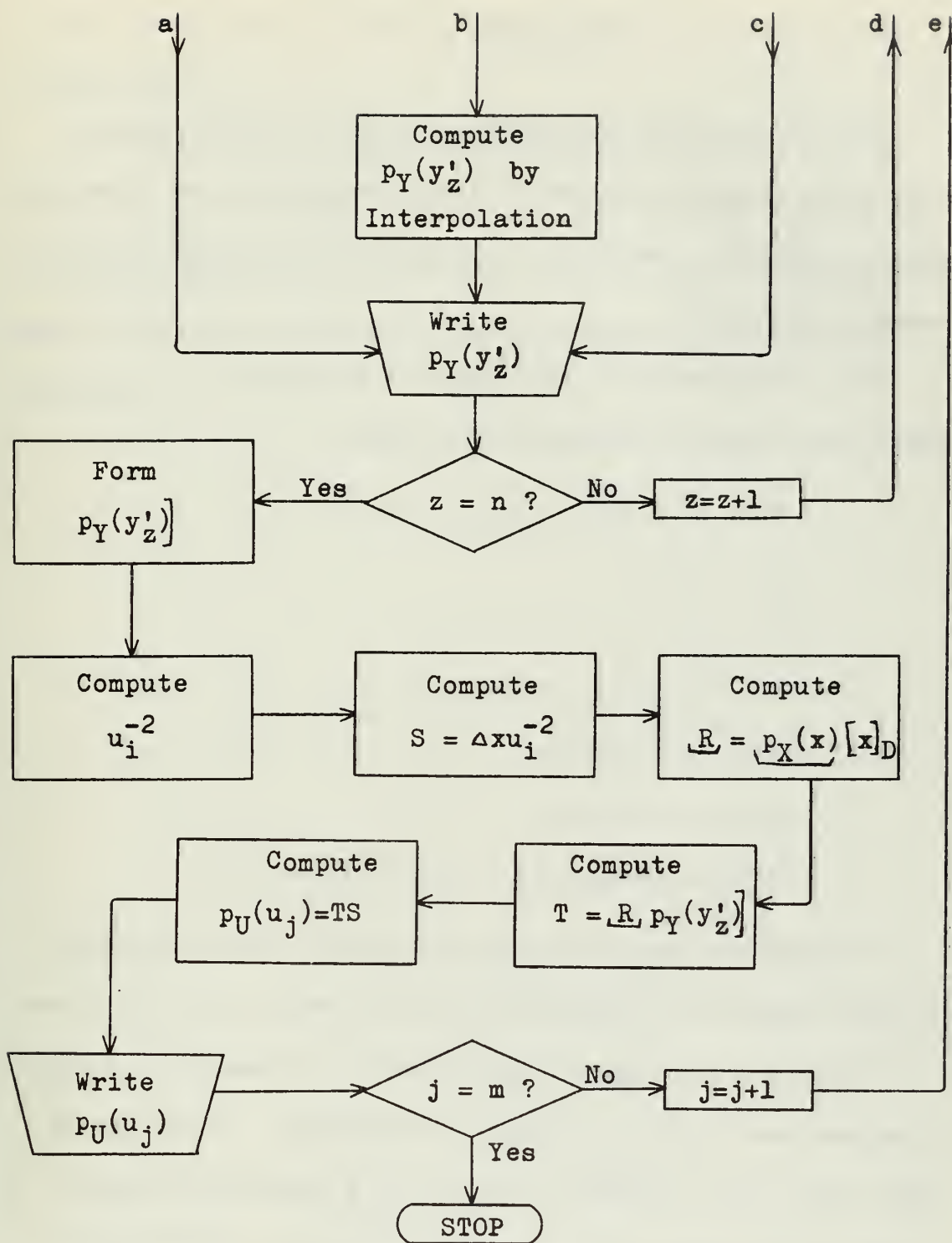


Figure 9

APPENDIX I

The solution to the deterministic problem has been obtained by using three separate programs. Linear interpolation and rectangular approximations have been used to perform all convolution and rectangular integration.

The problem and given information is as follows:

Obtain the cumulative distribution of θ where

$$\theta = C + \frac{A}{X(y)} + \frac{C}{Y(y)} + W(w)$$

$$C = 0.175$$

$$A = 8.1$$

$$B = 27.8$$

$$X(x) \text{ is } N(130;10)$$

$$Y(y) \text{ is } N(65;12)$$

$$W(w) \text{ is uniform } 0.875 \text{ from } -4/7 \text{ to } +4/7$$

Program one inverts the distribution x vs. $p(x)$ and multiplies it by the constant A . Program two does the same for $B/Y(y)$. Since both $X(x)$ and $Y(y)$ are normal distributions, the input data for these programs was taken from standard normal tables. 25 data points were used in each, the median and each $1/4$ standard deviation out to 3 standard deviations. The resulting distributions, $X'(x)$ and $Y'(y)$, are then used as input data for program three. Program three performs the convolution $U(u) = X'(x) + Y'(y)$, the addition $U'(u) = U(u) + C$,

and the convolution $\theta = U'(u) + W(w)$ to obtain the cumulative distribution of θ .

To obtain highly accurate results a Δx of 0.002 was used for convolution and rectangular integration. The over-all accuracy for this type of solution will depend upon how closely the rectangular approximations resemble the continuous distributions which they represent.

APPENDIX I - PROGRAM ONE

C EVALUATION OF THE PROBABILITY DENSITY FUNCTION, $P(X')$,
C WHERE $X' = A/X$ AND A , X , AND $P(X)$ ARE KNOWN. $P(X)$ HAS
C NORMAL DISTRIBUTION. N =NUMBER OF POINTS $X(I)$, M IS THE
C STANDARD DEVIATION OF THE DISTRIBUTION $P(X)$.

```

DIMENSION X(50),PX(50),U(50),PU(50)
READ(5,100)N
READ(5,101)M
READ(5,102)A
100 FORMAT(I2)
101 FORMAT(I2)
102 FORMAT(F10.0)
WRITE(6,210) N
210 FORMAT(1H1,T20,'N = ',I2//)
WRITE(6,211) M
211 FORMAT(T20,'M = ',I2//)
WRITE(6,212) A
212 FORMAT(T20,'A = ',F10.5//)
WRITE(6,213)
213 FORMAT(T20,'X(I)',T50,'PX(I)')//)
DO 300 I=1,N
  READ(5,103) X(I),PX(I)
103 FORMAT(2F10.0)
  WRITE(6,214) X(I),PX(I)
214 FORMAT(2F30.8)
300 CONTINUE
DO 301 I=1,N
  K=N-I+1
  PU(K)=(X(I)**2/A)*(PX(I)/M)
  U(K)=A/X(I)
301 CONTINUE
WRITE(6,200)
200 FORMAT(/,T20,'X',T50,'P(X )',/)
DO 302 I=1,N
201 FORMAT(2F30.8)
STOP
END

```

N = 25
M = 10
A = 8.1

X'	PX'
0.50624096E-01	0.13904164E 01
0.51428568E-01	0.27868738E 01
0.52258059E-01	0.51905861E 01
0.53114749E-01	0.91015186E 01
0.53999994E-01	0.14999995E 02
0.54915249E-01	0.23179779E 02
0.55862062E-01	0.33614029E 02
0.56842100E-01	0.45776794E 02
0.57857137E-01	0.58558014E 02
0.58909085E-01	0.70279877E 02
0.59999995E-01	0.79222473E 02
0.61132070E-01	0.83914819E 02
0.62307686E-01	0.83227264E 02
0.63529372E-01	0.77608536E 02
0.64799964E-01	0.67920517E 02
0.66122413E-01	0.55782471E 02
0.67499995E-01	0.43022202E 02
0.68936110E-01	0.31123703E 02
0.70434749E-01	0.21143661E 02
0.71999967E-01	0.13484363E 02
0.73636353E-01	0.80666637E 01
0.75348794E-01	0.45226297E 01
0.77142835E-01	0.23819437E 01
0.79024374E-01	0.11803312E 01
0.80999970E-01	0.54320967E 00

APPENDIX I - PROGRAM TWO

C EVALUATION OF THE PROBABILITY DENSITY FUNCTION, $P(Y')$,
C WHERE $Y'=B/Y$ AND B , Y , AND $P(Y)$ ARE KNOWN. $P(Y)$ HAS
C NORMAL DISTRIBUTION. N =NUMBER OF POINTS $Y(I)$, M IS THE
C STANDARD DEVIATION OF THE DISTRIBUTION $P(Y)$.

```

DIMENSION X(50),PX(50),U(50),PU(50)
READ(5,100)N
READ(5,101)M
READ(5,102)A
100 FORMAT(I2)
101 FORMAT(I2)
102 FORMAT(F10.0)
WRITE(6,210) N
210 FORMAT(1H1,T20,'N = ',I2//)
WRITE(6,211) M
211 FORMAT(T20,'M = ',I2//)
WRITE(6,212) A
212 FORMAT(T20,'A = ',F10.5//)
WRITE(6,213)
213 FORMAT(T20,'Y(I)',T50,'PY(I)')//)
DO 300 I=1,N
READ(5,103) X(I),PX(I)
103 FORMAT(2F10.0)
WRITE(6,214) X(I),PX(I)
214 FORMAT(2E30.8)
300 CONTINUE
DO 301 I=1,N
K=N-I+1
PU(K)=(X(I)**2/A)*(PX(I)/M)
U(K)=A/X(I)
301 CONTINUE
WRITE(6,200)
200 FORMAT(/,T20,'Y',T50,'P(Y )',//)
DO 302 I=1,N
201 FORMAT(2E30.8)
STOP
END

```

N = 25
M = 12
A = 27.8

Y'	PY'
0.27524740E 00	0.13454551E 00
0.28367329E 00	0.26197964E 00
0.29263145E 00	0.47343379E 00
0.30217373E 00	0.80428284E 00
0.31235939E 00	0.12821751E 01
0.32325566E 00	0.19132919E 01
0.33493960E 00	0.26742373E 01
0.34749985E 00	0.35031176E 01
0.36103874E 00	0.43010139E 01
0.37567550E 00	0.49425173E 01
0.39154911E 00	0.53205528E 01
0.40882331E 00	0.53600159E 01
0.42769212E 00	0.50520163E 01
0.44838685E 00	0.44558611E 01
0.47118622E 00	0.36740417E 01
0.49642831E 00	0.28304844E 01
0.52452803E 00	0.20377035E 01
0.55599976E 00	0.13684053E 01
0.59148908E 00	0.85751051E 00
0.63181788E 00	0.50082946E 00
0.67804843E 00	0.27210426E 00
0.73157859E 00	0.13721460E 00
0.79428536E 00	0.64261079E-01
0.86874962E 00	0.27932845E-01
0.95862025E 00	0.11092324E-01

APPENDIX I - PROGRAM THREE

CALCULATION OF $U=X+Y$ WHERE X , $P(X)$, Y , $P(Y)$, ARE GIVEN AND THE GIVEN VALUES OF $X(I)$ AND $Y(I)$ ARE INCREMENTED BY DX .

```

DIMENSION X(500),PX(500),Y(500),PY(500),U(500),PU(500)
PW(600),Z(1000),PZ(1000),F(500),W(600)
READ(5,101) DX
101 FORMAT(F10.0)
WRITE(6,210) DX
210 FORMAT(1H1,T20,'DX =',F10.3//)
READ(5,106) C
106 FORMAT(F10.0)
WRITE(6,213) C
213 FORMAT(T20,'C =',F10.8//)
READ(5,105) N
105 FORMAT(I5)
WRITE(6,211) N
211 FORMAT(T20,'N =',I5//)
READ(5,104) M
104 FORMAT(I5)
WRITE(6,212) M
212 FORMAT(//,T20,'M =',I5//)
WRITE(6,202)
202 FORMAT(//,T20,'X(I)',T50,'PX(I)')//)
CALL ONE(X,PX)
WRITE(6,203)
203 FORMAT(//,T20,'Y(I)',T50,'PY(I)')//)
CALL ONE(Y,PY)
K=N+M-1
U(1)=X(1)+Y(1)
DO 302 I=2,K
302 U(I)=U(I-1)+DX
DO 303 J=1,K
PU(J)=0.0
DO 304 I=1,J
IF((J-I+1).GT.M) GO TO 304
IF(I.GT.N) GO TO 303
PU(J)=PU(J)+PX(I)*PY(J-I+1)*DX
304 CONTINUE
303 CONTINUE

```

C CALCULATION OF $U(J)=U(J)+C$ WHERE 'C' IS A CONSTANT

```

DO 306 I=1,K
U(I)=U(I)+C
306 CONTINUE
WRITE(6,200)
200 FORMAT(1H1,T20,'U(I)',T50,'PU(I)')//)
DO 305 I=1,K
305 WRITE(6,201) U(I),PU(I)
201 FORMAT(2E30.8)

```

C CALCULATION AND PLOT OF CUMULATIVE DISTRIBUTION

```

F(1)=PU(1)*DX
DO 320 I=2,K
F(I)=F(I-1)+PU(I)*DX
320 CONTINUE
WRITE(6,225)
225 FORMAT(1H1,T20,'U(J)',T40,'CUMULATIVE DISTRIBUTION')
DO 321 I=1,K
WRITE(6,201) U(I),F(I)
321 CONTINUE

```

PERFORM THE CONVOLUTION OF 'U' WITH THE CONSTANT PDF 'W' ($Z=U+W$).

```

CALL TWO(W,PW,L)
NM=K+L-1

```

```

      Z(1)=U(1)+W(1)
      DO 310 I=2,NM
310  Z(I)=Z(I-1)+DX
      DO 311 J=1,NM
      PZ(J)=0.0
      DO 312 I=1,J
      IF((J-I+1).GT.L) GO TO 312
      IF(I.GT.K) GO TO 311
      PZ(J)=PZ(J)+PU(I)*PW(J-I+1)*DX
312  CONTINUE
311  CONTINUE
      WRITE(6,230)
230  FORMAT(1H1,T20,'Z(I)',T50,'PZ(I)')//)
      DO 313 I=1,NM
313  WRITE(6,231) Z(I),PZ(I)
231  FORMAT(2E30.8)

```

C CALCULATION OF THE CUMULATIVE DISTRIBUTION OF 'Z'.

```

      PZ(1)=PZ(1)*DX
      DO 330 I=2,NM
      PZ(I)=PZ(I-1)+PZ(I)*DX
330  CONTINUE
      WRITE(6,220)
220  FORMAT(1H1,T20,'U(I)',T40,'CUMULATIVE DISTRIBUTION')
      DO 331 I=1,NM
221  FORMAT(2E30.8)
331  CONTINUE
      STOP
      END

```

SUBROUTINE ONE(XE,PXE)

C FROM GIVEN VALUES OF X(I) AND PX(I), (THESE ARE U(I) AND PU(I) PREVIOUSLY CALCULATED), AND A CHOSEN DX, WE OBTAIN THE VALUES OF X(J), WITH INTERVALS DX, AND CALCULATE THE CORRESPONDING VALUES OF PX(J) BY LINEAR INTERPOLATION.

```

      DIMENSION X(500),PX(500),XE(500),PXE(500)
      READ(5,100) N
      FORMAT(I2)
100  READ(5,101) DX
101  FORMAT(F10.0)
      DO 300 I=1,N
300  READ(5,102) X(I), PX(I)
102  FORMAT(2F10.0)
      I=1
      J=1
      XE(J)=X(I)+DX/2
400  IF(XE(J)-X(I)) 500,501,502
502  I=I+1
      IF(I.LT.(N+1)) GO TO 400
      GO TO 203
501  PXE(J)=PX(I)
      GO TO 200
500  PXE(J)=PX(I-1)+(PX(I)-PX(I-1))/(X(I)-X(I-1))*(XE(J)-X(I-1))
200  WRITE(6,201) XE(J), PXE(J)
201  FORMAT(2E30.8)
      XE(J+1)=XE(J)+DX
      J=J+1
      GO TO 400
203  RETURN
      END

```

SUBROUTINE TWO(W,PW,L)

C THIS PROGRAM TAKES THE GIVEN CONSTANT PROBABILITY
 C FUNCTION 'W', WITH AMPLITUDE 0.875 FROM -4/7 TO +4/7,

AND OBTAINS THE VALUES OF W(I) WITH INCREMENT DX=0.002

```
DIMENSION W(600), PW(600)
W(1)=-0.571
PW(1)=0.875
DX=0.002
L=571
WRITE(6,250)
250 FORMAT(1H1,T20,'W(I)',T50,'PW(I)')
DO 350 I=2,L
W(I)=W(I-1)+DX
PW(I)=PW(I-1)
350 CONTINUE
DO 351 I=1,L
WRITE(6,251) W(I),PW(I)
251 FORMAT(2E30.8)
351 CONTINUE
RETURN
END
```

```
N = 15
M = 242
C = 0.175
DX = 0.002
```

THETA

CUMULATIVE DISTRIBUTION

-0.63128109E-01	0.351622280E-08
-0.65128075E-01	0.24354559E-07
-0.64128041E-01	0.93342241E-07
-0.62128045E-01	0.29366913E-06
-0.60128044E-01	0.71189930E-06
-0.58128044E-01	0.14791967E-05
-0.56128044E-01	0.27345241E-05
-0.54128043E-01	0.46203226E-05
-0.52128043E-01	0.72799021E-05
-0.50128043E-01	0.10858982E-04
-0.48128042E-01	0.15508529E-04
-0.46128042E-01	0.21389027E-04
-0.44128042E-01	0.28669005E-04
-0.42128041E-01	0.37537306E-04
-0.40128041E-01	0.48194212E-04
-0.38128041E-01	0.60856095E-04
-0.36128040E-01	0.75755495E-04
-0.34128040E-01	0.93142065E-04
-0.32128040E-01	0.11328337E-03
-0.30128039E-01	0.13646450E-03
-0.28128039E-01	0.16298697E-03
-0.26128039E-01	0.19316789E-03
-0.24128038E-01	0.22734021E-03
-0.22128038E-01	0.25585301E-03
-0.20128038E-01	0.30907127E-03
-0.18128037E-01	0.35737362E-03
-0.16128037E-01	0.41115121E-03
-0.14128037E-01	0.47080661E-03
-0.12128036E-01	0.53675333E-03
-0.10128036E-01	0.60941558E-03
-0.81280358E-02	0.68922667E-03
-0.61280355E-02	0.77662780E-03
-0.41280352E-02	0.87206601E-03
-0.21280353E-02	0.97599253E-03
-0.12803543E-03	0.10888616E-02
0.13719644E-02	0.12111294E-02
0.38719643E-02	0.13432535E-02
0.53719628E-02	0.14856923E-02
0.78719594E-02	0.16389030E-02
0.98719560E-02	0.18033390E-02
0.11871953E-01	0.19794477E-02
0.13871949E-01	0.21676687E-02
0.15871946E-01	0.23684341E-02
0.17871942E-01	0.25821684E-02
0.19871939E-01	0.28092891E-02
0.21871936E-01	0.30502048E-02
0.23871932E-01	0.33053122E-02
0.25871929E-01	0.35749928E-02
0.27871925E-01	0.38596098E-02
0.29871922E-01	0.41595064E-02
0.31871919E-01	0.447750199E-02
0.33871915E-01	0.48064664E-02
0.35871912E-01	0.51541552E-02
0.37871908E-01	0.55183806E-02
0.39871905E-01	0.589994144E-02
0.41871902E-01	0.62975064E-02
0.43871898E-01	0.67128837E-02
0.45871895E-01	0.71457513E-02
0.47871891E-01	0.75962991E-02
0.49871888E-01	0.80647096E-02
0.51871885E-01	0.85511431E-02
0.53871881E-01	0.90557523E-02
0.55871878E-01	0.95786639E-02
0.57871874E-01	0.10119986E-01
0.59871871E-01	0.10679796E-01
0.61871868E-01	0.11258151E-01
0.63871861E-01	0.11855092E-01
0.65871835E-01	0.12470651E-01
0.67871809E-01	0.13104849E-01

C.69871783E-01
 O.71871758E-01
 O.73871732E-01
 C.75871706E-01
 C.77871680E-01
 C.79871655E-01
 O.81871629E-01
 C.83871603E-01
 O.85871577E-01
 C.87871552E-01
 C.89871526E-01
 C.91871500E-01
 C.93871474E-01
 C.95871449E-01
 O.97871423E-01
 C.99871397E-01
 C.10187137E 00
 C.10387135E 00
 C.10587132E 00
 C.10787129E 00
 C.10987127E 00
 C.11187124E 00
 C.11387122E 00
 C.11587119E 00
 C.11787117E 00
 C.11987114E 00
 C.12187111E 00
 C.12387109E 00
 C.12587106E 00
 C.12787104E 00
 C.12987101E 00
 C.13187099E 00
 C.13387096E 00
 C.13587093E 00
 C.13787091E 00
 C.13987088E 00
 C.14187086E 00
 C.14387083E 00
 C.14587080E 00
 C.14787078E 00
 C.14987075E 00
 C.15187073E 00
 C.15387070E 00
 C.15587068E 00
 C.15787065E 00
 C.15987062E 00
 C.16187060E 00
 C.16387057E 00
 C.16587055E 00
 C.16787052E 00
 C.16987050E 00
 C.17187047E 00
 C.17387044E 00
 C.17587042E 00
 C.17787039E 00
 C.17987037E 00
 C.18187034E 00
 C.18387032E 00
 C.18587029E 00
 C.18787026E 00
 C.18987024E 00
 C.19187021E 00
 C.19387019E 00
 C.19587016E 00
 C.19787014E 00
 C.19987011E 00
 C.20187008E 00
 C.20387006E 00
 C.20587003E 00
 C.20787001E 00
 C.20986998E 00
 C.21186996E 00

O.13757691E-01
 O.14429174E-01
 O.15119277E-01
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 O.16555130E-01
 O.17300732E-01
 O.18064659E-01
 O.18846810E-01
 O.19647073E-01
 O.20465329E-01
 O.21301460E-01
 O.22155326E-01
 O.23026783E-01
 O.23915663E-01
 O.24821792E-01
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 O.26685044E-01
 O.27541781E-01
 O.28614990E-01
 O.29604476E-01
 O.30510032E-01
 O.31631447E-01
 O.32668509E-01
 O.33720996E-01
 O.34798683E-01
 O.35871338E-01
 O.36968727E-01
 O.38080610E-01
 O.39206754E-01
 O.40346917E-01
 O.41500859E-01
 O.42668343E-01
 O.43849129E-01
 O.45042980E-01
 O.46249662E-01
 O.47468934E-01
 O.48700567E-01
 O.49944323E-01
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 O.52467272E-01
 O.53745996E-01
 O.55035912E-01
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 O.79917431E-01
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 O.82848191E-01
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 O.85806727E-01
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 O.88791728E-01
 O.90293765E-01
 O.91802001E-01
 O.93316257E-01
 O.94836414E-01
 O.96362293E-01
 O.97893715E-01

C.21386993E 00
 O.21586990E 00
 O.21786988E 00
 O.21986985E 00
 O.22186983E 00
 O.22386980E 00
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APPENDIX II

The tabulated sample data for the non-deterministic case, Table II, was obtained from a modified form of equation (19). Using the modified equation,

$$\phi = 0.175 + \frac{27.8}{Y(y)} + \frac{8.1}{X(x)}$$

where $X(x)$ is $N(130;10)$

$Y(y)$ is $N(65;12)$

values of ϕ and sensitivity were calculated for chosen values of x and y . Assuming these to be measured values, the data was then used to obtain the cumulative distribution of ϕ .

Since the data indicates ϕ is more sensitive to changes in y than to changes in x , the problem is worked with x as a constant and y as a single variable as discussed in section III. Aitkin-Hermite interpolation [Hildebrand 1956] is used to estimate the values of y between data points. This interpolation scheme utilizes the sensitivity data and obtains a more accurate estimate than would linear interpolation. The calculation of the cumulative distribution of ϕ is then made using equation (35) in discrete form. As in the deterministic case, a Δx of 0.002 was used for rectangular integration.

APPENDIX II - PROGRAM ONE

```

C      CALCULATION OF THE CUMULATIVE DISTRIBUTION OF AN
      OUTPUT, PHI, FROM MEASURED VALUES OF THE PARAMETERS
      X, Y, AND THE SENSITIVITY OF PHI WITH RESPECT TO X
      AND Y. THE PROBABILITY DISTRIBUTIONS OF X AND Y ARE
      KNOWN, N(130;10) AND N(65;12) RESPECTIVELY, AND THE
      DATA POINTS HAVE BEEN SELECTED IN DX INTERVALS TO
      FACILITATE THE CALCULATIONS.

      DIMENSION ARG1(5,5),VAL(10),ARG(5),XC(5),F(5,2),Z(5),
      PYT(25),PXC(5),X1(200),PX(200),YT(25)

C      READ THE MEASURED VALUES OF X AND THEIR CORRESPONDING
C      PROBABILITIES, DENOTED BY XC AND PXC.

      READ(5,102) (XC(I),PXC(I),I=1,5)
      WRITE(6,103)
103  FORMAT(//,T21,'XC',T47,'PXC',/)
      WRITE(6,106) (XC(I),PXC(I),I=1,5)
106  FORMAT(2F26.5)
101  FORMAT(F10.5)

      READ THE MEASURED VALUES OF PHI, DENOTED BY ARG1.
      THERE IS A POINT PHI CORRESPONDING TO EACH SET OF
      X - Y POINTS.

      READ(5,101) ((ARG1(I,J),J=1,5),I=1,5)

C      READ THE MEASURED VALUES OF Y AND THE CORRESPONDING
      SENSITIVITY OF PHI WITH RESPECT TO Y FOR EACH OF
      THESE VALUES, DENOTED BY F(I,J).

      READ(5,102) ((F(I,J),J=1,2),I=1,5)
102  FORMAT(2F10.5)

      READ THE KNOWN CUMULATIVE DISTRIBUTION OF Y, DENOTED
      BY YT AND PYT.

      READ(5,102) (YT(I),PYT(I),I=1,25)
      X1(1)=0.5708710
      DC 204 K=1,168
      PT=0.0

C      X DENOTES THE VALUE OF PHI FOR WHICH THE CUMULATIVE
C      DISTRIBUTION IS BEING CALCULATED.

      X=X1(K)
      DC 202 I=1,5
      DC 201 J=1,5

      Z(J) DENOTES THE ARRAY OF ARG1 VALUES CORRESPONDING
      TO A PARTICULAR VALUE OF XC.

201  Z(J)=ARG1(I,J)
      DX=10.0
      EPS=0.001
      NDIM=5

C      SUBROUTINE ATSM ORDERS THE ARRAY Z(J) SUCH THAT
      ABS(ARG(I)-X).GE.ABS(ARG(J)-X) IF I.GE.J. THIS IS
      REQUIRED FOR SUBROUTINE ONE.

      CALL ATSM(X,Z,F,5,2,ARG,VAL,NDIM)

C      SUBROUTINE ONE INTERPOLATES FUNCTION VALUE Y FOR A
      GIVEN ARGUMENT VALUE X USING THE TABLE (ARG,VAL) OF
      ARGUMENT, FUNCTION, AND DERIVATIVE VALUES FROM THE
      SUBROUTINE ATSM. THIS SUBROUTINE USES THE AITKEN-
      HERMITE INTERPOLATION.

```



```

      CALL ONE(X,ARG,VAL,Y,NDIM,EPS,IER)
C      SUBROUTINE TWO USES LINEAR INTERPOLATION TO OBTAIN
      P(Y) FOR THE VALUE OF Y OBTAINED IN SUBROUTINE ONE.
      CALL TWO(Y,PY,YT,PYT)
      PT=PT+(1-PY)*PXC(I)
      PX(K)=PT*DX
302  FORMAT(2E30.8)
202  CONTINUE
      X1(K+1)=X1(K)+0.0020000
204  CONTINUE
      WRITE(6,107)
107  FORMAT(1H1,T20,'U(I)',T42,'CUMULATIVE DISTRIBUTION',//
      WRITE(6,302) (X1(K),PX(K),K=1,168)
      STOP
      END

      SUBROUTINE ATSM(X,Z,F,IROW,ICOL,ARG,VAL,NDIM)
      DIMENSION Z(1),F(1),ARG(1),VAL(1)
C      CASE IROW=1 IS CHECKED OUT
      IF(IROW-1)23,21,1
1      N=NDIM
C      IF N IS GREATER THAN IROW, N IS SET EQUAL TO IROW.
      IF(N-IROW)3,3,2
2      N=IROW
C      CASE IROW .GE. 2
C      SEARCHING FOR SUBSCRIPT J SUCH THAT Z(J) IS NEXT TO X.
3      IF(Z(IROW)-Z(1))5,4,4
4      J=IROW
      I=1
      GO TO 6
5      I=IROW
      J=1
      K=(J+I)/2
      IF(X-Z(K))7,7,8
7      J=K
      GO TO 9
8      I=K
9      IF(ABS(J-I)-1)10,10,6
10     IF(ABS(Z(J)-X)-ABS(Z(I)-X))12,12,11
11     J=I
C      TABLE SELECTION
12     K=J
      JL=0
      JR=0
      DO 20 I=1,N
      AFG(I)=Z(K)
      IF(ICOL-1)14,14,13
13     VAL(2*I-1)=F(K)
      KK=K+IROW
      VAL(2*I)=F(KK)
      GO TO 15
14     VAL(I)=F(K)
15     JJR=J+JR
      IF(JJR-IROW)16,18,18
16     JJL=J-JL
      IF(JJL-1)19,19,17
17     IF(ABS(Z(JJR+1)-X)-ABS(Z(JJL-1)-X))19,19,18
18     JL=JL+1
      K=J-JL
      GO TO 20
19     JR=JR+1
      K=J+JR
20     CONTINUE
      RETURN

```

```

C      CASE IFOW=1
21     ARG(1)=Z(1)
      VAL(1)=F(1)
      IF(ICOL-2)23,22,23
22     VAL(2)=F(2)
23     RETURN
      END

      SUBROUTINE ONE(X,ARG,VAL,Y,NDIM,EPS,IER)
      DIMENSION ARG(1),VAL(1)
      IER=2
      H2=X-ARG(1)
      IF(NDIM-1)2,1,3
1     Y=VAL(1)+VAL(2)*H2
2     RETURN

C      VECTOR ARG HAS MORE THAN 1 ELEMENT.
C      THE FIRST STEP PREPARES VECTOR VAL SUCH THAT AITKEN
C      SCHEME CAN BE USED.
3     I=1
      DO 5 J=2,NDIM
      H1=H2
      H2=X-ARG(J)
      Y=VAL(I)
      VAL(I)=Y+VAL(I+1)*H1
      H=H1-H2
      IF(H)4,13,4
4     VAL(I+1)=Y+(VAL(I+2)-Y)*H1/H
5     I=I+2
      VAL(I)=VAL(I)+VAL(I+1)*H2
C      END OF FIRST STEP

C      PREPARE AITKEN SCHEME
      DELT2=0.
      IEND=I-1

C      START AITKEN-LOOP
      DO 9 I=1,IEND
      DELT1=DELT2
      Y=VAL(1)
      M=(I+3)/2
      H1=ARG(M)
      DO 6 J=1,I
      K=I+1-J
      L=(K+1)/2
      H=ARG(L)-H1
      IF(H)6,14,6
6     VAL(K)=(VAL(K)*(X-H1)-VAL(K+1)*(X-ARG(L)))/H
      DELT2=ABS(Y-VAL(1))
      IF(DELT2-EPS)11,11,7
7     IF(I-5)9,8,3
8     IF(DELT2-DELT1)9,12,12
9     CONTINUE
C      END OF AITKEN-LOOP

10    Y=VAL(1)
      RETURN

      THERE IS SUFFICIENT ACCURACY WITHIN 2*NDIM-2 ITERATION
      STEPS
11    IEP=0
      GOTO 10

C      TEST VALUE DELT2 STARTS OSCILLATING
12    IER=1
      RETURN

C      THERE ARE TWO IDENTICAL ARGUMENT VALUES IN VECTOR ARG
13    Y=VAL(1)
14    IER=3

```


RETURN
END

SUBROUTINE TWO(Y,PY,YT,PYT)

DIMENSION YT(25),PYT(25)
IF(Y.LT.YT(1)) GO TO 301
J=2
110 IF(Y-YT(J)) 200,201,202
202 J=J+1
IF(J.LE.25) GO TO 110
PY=1.0
GO TO 300
201 PY=PYT(J)
GO TO 300
200 PY=PYT(J-1)+(PYT(J)-PYT(J-1))/(YT(J)-YT(J-1))*(Y-YT(J-1))
GO TO 300
301 PY=C.0
300 RETURN
END

PHI

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 0.57487094E 00
 0.57687092E 00
 0.57887089E 00
 0.58087087E 00
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 0.58487082E 00
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13. ABSTRACT <p>Techniques for extrapolating statistical data on the parameters of individual components to the statistical performance index for an overall system are considered. Two cases are evaluated. (1) The deterministic case in which the system's performance index is known functionally in terms of the system parameters. (2) The non-deterministic case in which only limited data on the performance index and its sensitivity with respect to system parameters is known. Computer programs are developed in both cases for combining given probability density distributions for the parameters into an overall probability density distribution for the system performance index. Theory and programs are developed and verified with a specific numerical example.</p>			

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Deterministic

Jacobian

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